

Phase-type Distributions in Healthcare Modelling II

Mark Fackrell

Department of Mathematics and Statistics
 University of Melbourne
 Victoria 3010
 AUSTRALIA

email: mfackrel@ms.unimelb.edu.au

In the second of these three articles on phase-type (*PH*) distributions in healthcare modelling I will introduce Coxian distributions, a particularly useful subclass, and explain briefly how they have been used in healthcare modelling. In the first article (see Fackrell [3]) the case for using stochastic models in healthcare was made, and *PH* distributions were introduced. Fackrell [2] contains a more comprehensive treatment and bibliography.

An order p Coxian distribution has a representation of the form

$$\boldsymbol{\alpha} = (\alpha_1 \quad \alpha_2 \quad \dots \quad \alpha_p) \quad (1)$$

$$\mathbf{T} = \begin{pmatrix} -\lambda_1 & \lambda_1 & 0 & \dots & 0 \\ 0 & -\lambda_2 & \lambda_2 & \dots & 0 \\ 0 & 0 & -\lambda_3 & \ddots & 0 \\ \vdots & \vdots & \ddots & \ddots & \vdots \\ 0 & 0 & 0 & \dots & -\lambda_p \end{pmatrix}, \quad (2)$$

where $0 < \lambda_1 \leq \lambda_2 \leq \dots \leq \lambda_p$. Figure 1 shows the state transition diagram for an order p Coxian distribution.

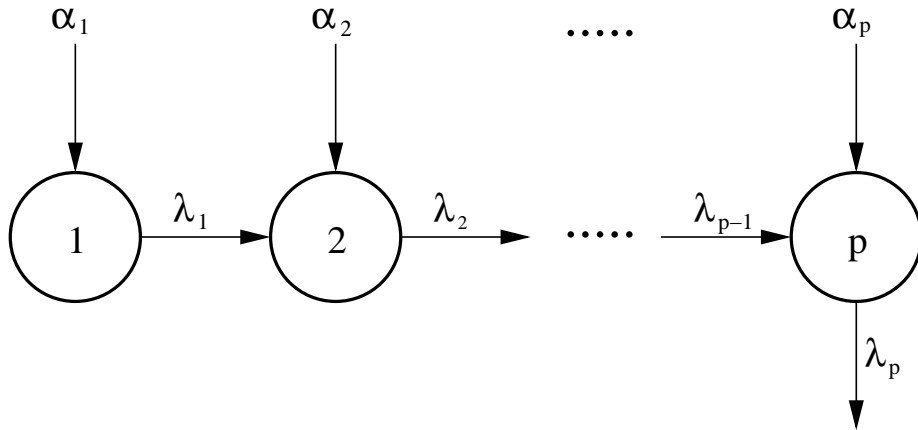


Figure 1: State transition diagram for an order p Coxian distribution.

The first thing we notice is that the representation (1)–(2) depends on only $2p - 1$ parameters. General *PH* representations require $p^2 + p - 1$ parameters, but only $2p - 1$ parameters are needed to define the distribution uniquely. Thus, general *PH* distributions are considerably overparameterized, whereas Coxian distributions are not. When it comes to fitting data with *PH* distributions, Coxian distributions have been the preferred option

for practitioners in healthcare modelling and elsewhere, partly because of this reason - there are a lot less parameters to estimate.

Coxian distributions exhibit quite a lot of flexibility. Figure 2 shows the density functions for three order 4 Coxian distributions with the same generator \mathbf{T} , but different vectors $\boldsymbol{\alpha}$. We can see by their shapes that they exhibit much more flexibility than the exponential distribution, which should come as no surprise - they rely on 7 free parameters instead of just one! Indeed, Coxian distributions are more versatile than both hyperexponential distributions and generalized Erlang distributions, which also depend on $2p-1$ parameters, see Fackrell [2].

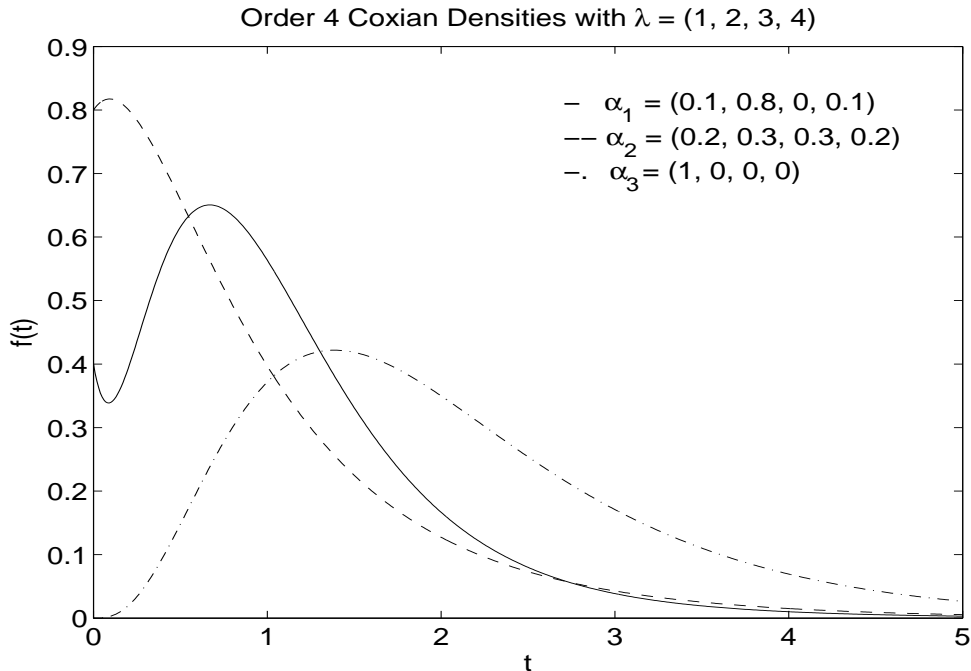


Figure 2: Density functions for three different order 4 Coxian distributions with $\lambda_1 = 1$, $\lambda_2 = 2$, $\lambda_3 = 3$, $\lambda_4 = 4$.

A curious mathematical fact is that any *PH* distribution whose generator \mathbf{T} is an upper triangular matrix, has a Coxian representation of the same or lower order, see Cumani [1] or O’Cinneide [9]. For example, the hyperexponential distribution with representation

$$\boldsymbol{\alpha} = \left(\frac{1}{6} \quad \frac{1}{2} \quad \frac{1}{3} \right) \quad (3)$$

$$\mathbf{T} = \begin{pmatrix} -2 & 0 & 0 \\ 0 & -3 & 0 \\ 0 & 0 & -1 \end{pmatrix}, \quad (4)$$

has an equivalent Coxian representation

$$\boldsymbol{\beta} = \left(\frac{1}{9} \quad \frac{1}{6} \quad \frac{13}{18} \right) \quad (5)$$

$$\mathbf{S} = \begin{pmatrix} -1 & 1 & 0 \\ 0 & -2 & 2 \\ 0 & 0 & -3 \end{pmatrix}, \quad (6)$$

In fact, any *PH* distribution whose generator \mathbf{T} has only real eigenvalues has a Coxian representation of *some* order. The million dollar question is, of course, how big is the order? It is widely believed that there are examples of *PH* distributions of relatively low order that have Coxian representations of high order. Nevertheless, Coxian distributions are a very important subclass of *PH* distributions and need to be studied.

Coxian distributions have been popular with healthcare modellers because the states (or groups of states) can sometimes be given a physical interpretation. For example, Xie, Chausalet, and Millard [10] modelled the length of stay (LOS) of geriatric patients in residential and nursing home care with two, 2-state Coxian distributions. Figure 3 shows the state transition diagram for their model. Here, patients enter the system via the

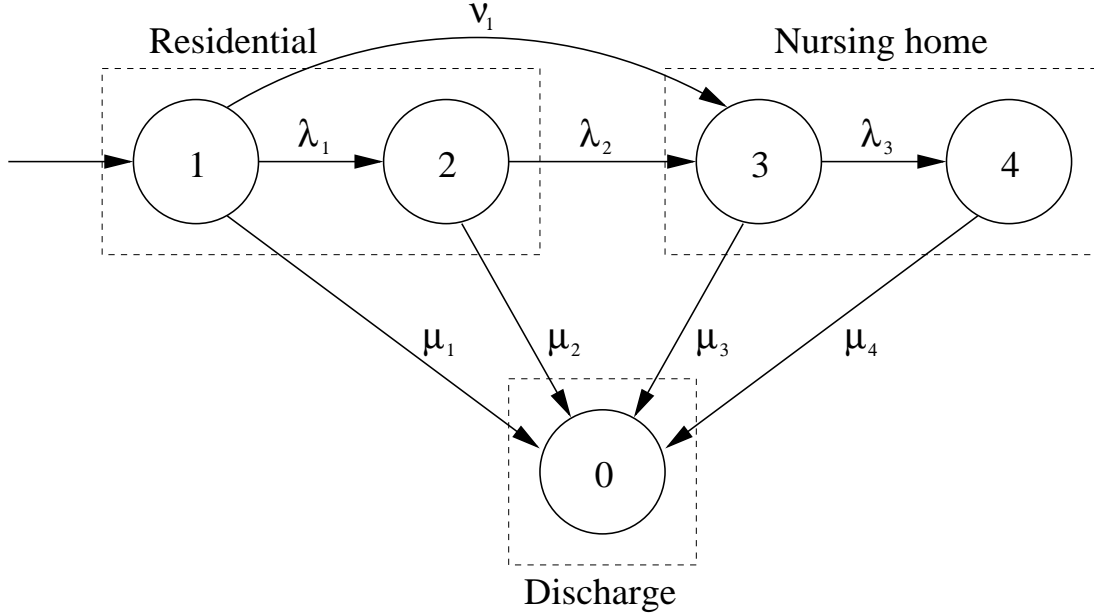


Figure 3: State transition diagram to model the length of stay in residential and nursing home care.

residential home care block where they can spend a short time (state 1 only), or a long time (state 1 followed by state 2). They can be discharged from either state, or progress to nursing home care, where again they can spend a short time (state 3), or a long time (states 3 and 4) before being discharged or dying.

The corresponding *PH* representation for the model is

$$\boldsymbol{\alpha} = (1 \ 0 \ 0 \ 0) \tag{7}$$

$$\mathbf{T} = \begin{pmatrix} -(\lambda_1 + \mu_1 + \nu_1) & \lambda_1 & \nu_1 & 0 \\ 0 & -(\lambda_2 + \mu_2) & \lambda_2 & 0 \\ 0 & 0 & -(\lambda_3 + \mu_3) & \lambda_3 \\ 0 & 0 & 0 & -\mu_4 \end{pmatrix}. \tag{8}$$

We remark here that because \mathbf{T} is an upper triangular matrix, the *PH* distribution with representation $(\boldsymbol{\alpha}, \mathbf{T})$ is a Coxian distribution.

The authors fitted the Coxian distribution to four years data from the social services department of a London borough using maximum likelihood estimation. They reported

$\lambda_1 = \lambda_2 = \mu_2 = 0$, $\nu_1 = 0.000228$, $\mu_1 = 0.000855$, $\lambda_3 = 0.010874$, $\mu_3 = 0.006138$, and $\mu_4 = 0.001275$. The LOS in residential care was modelled by an exponential distribution (state 2 was unnecessary) with an average LOS of 923 days, with 21% of patients moving on to nursing home care and the rest getting discharged. The LOS for short stay patients in nursing home care (36 %) was modelled with an exponential distribution (average LOS 59 days), and the LOS for long stay patients (64 %) was modelled with a 2-state generalized Erlang distribution (average LOS 843 days).

Other notable papers where Coxian distributions have been used to model systems in healthcare include Faddy and McClean [4] and [5], Faddy and Taylor [6], McClean, Faddy, and Millard [8], and Marshall and McClean [7].

In the next issue's article I will propose some new ways in which *PH* distributions could be used in healthcare modelling.

References

- [1] Cumani, A. On the canonical representation of homogeneous Markov processes modelling failure-time distributions. *Microelectronics and Reliability* **1982**, *22*, 583–602.
- [2] Fackrell, M. Using phase-type distributions in health and social care modelling. Submitted, 2007.
- [3] Fackrell, M. Phase-type distributions in health care modelling I. *Nosokinetic News*, Issue 4.4, August, 2007.
- [4] Faddy, M. J.; McClean, S. I. Analysing data on lengths of stay of hospital patients using phase-type distributions. *Applied Stochastic Models in Business and Industry* **1999**, *15*, 311–317.
- [5] Faddy, M. J.; McClean, S. I. Markov chain modelling for geriatric patient care. *Methods of Information in Medicine* **2005**, *44*, 369–373.
- [6] Faddy, M. J.; Taylor, G. T. Stochastic modelling of the onset of bronchiolitis obliterans syndrome following lung transplantations: An analysis of risk factors. *Mathematical and Computer Modelling* **2003**, *38*, 1185–1189.
- [7] Marshall, A. H.; McClean, S. I. Conditional phase-type distributions for modelling length of stay in hospital. *International Transactions in Operational Research* **2003**, *10*, 565–576.
- [8] McClean, S. I.; Faddy, M. J.; Millard, P. H. Markov model-based clustering for efficient patient care. *Proceedings of the 18th IEEE Symposium on Computer-Based Medical Systems* **2005**, 467–472.
- [9] O’Cinneide, C. A. On non-uniqueness of representations of phase-type distributions. *Communications in Statistics – Stochastic Models* **1989**, *5*, 247–259.
- [10] Xie, H.; Chausalet, T. J.; Millard, P. H. A continuous time Markov model for the length of stay of elderly people in institutional long-term care. *Journal of the Royal Statistical Society. Series A* **2005**, *168*, 51–61.